

Online Supplement to “Common Bubble Detection in Large Dimensional Financial Systems”^{*}

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This document is an Online Supplement to Chen, Phillips & Shi (2019). Section A provides the proofs of Lemma 4.1 in Chen, Phillips & Shi (2019) which states the consistency of the estimated factor under the null. The arguments follow directly from Bai (2004) and Chen, Li & Phillips (2019). Section B presents the Monte Carlo simulation results for the real time PSY-factor procedure.

A Proof of Lemma 4.1

We introduce the main notation again for ease of reference. The data generating process under the null is

$$X = F^0 \Lambda^{0'} + E, \tag{1}$$

where $X = (\underline{X}_1 \dots, \underline{X}_N)$ is a $T \times N$ matrix of the observed data with $\underline{X}_i = (X_{i1}, \dots, X_{iT})'$, $F^0 = (f_{0,1} \dots, f_{0,T})'$ is a $T \times 1$ vector, $\Lambda^0 = (\lambda_{0,1} \dots, \lambda_{0,N})'$ is a $N \times 1$ vector of loading coefficients, and $E = (\underline{e}_1 \dots, \underline{e}_N)$ is a $T \times N$ matrix of idiosyncratic errors with $\underline{e}_i = (e_{i1}, \dots, e_{iT})'$. Alternatively,

$$X_t = \Lambda^0 f_{0,t} + e_t, \tag{2}$$

where $X_t = (X_{1t}, \dots, X_{Nt})'$ and $e_t = (e_{1t}, \dots, e_{Nt})'$.

The common factor F^0 and the factor loadings Λ^0 are estimated by the principal component analysis, subject to a normalization on the loadings. The estimated loadings for the first factor is denoted by $\tilde{L}_1 = (\tilde{l}_{1,1}, \tilde{l}_{1,2}, \dots, \tilde{l}_{1,N})'$, which is \sqrt{N} times the eigenvectors corresponding to the largest eigenvalue (denoted by λ) of the $N \times N$ matrix $X'X$. The estimated first common factor is denoted by \tilde{y} and defined as

$$\tilde{y} = X \tilde{L}_1 / N.$$

Let λ_{NT} be the largest eigenvalues of $\frac{1}{NT^2} X'X$. By the definition of eigenvectors and eigenvalues, we have the mathematical equality

$$\frac{1}{NT^2} X'X \tilde{L}_1 = \tilde{L}_1 \lambda_{NT}. \tag{3}$$

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It follows that

$$\tilde{L}_1 = \frac{1}{NT^2} X' X \tilde{L}_1 \lambda_{NT}^{-1} = \frac{1}{NT^2} (\Lambda^0 F^{0'} F^0 \Lambda^{0'} + \Lambda^0 F^{0'} E + E' F^0 \Lambda^{0'} + E' E) \tilde{L}_1 \lambda_{NT}^{-1} \quad (4)$$

using (1). Let $H = \lambda_{NT}^{-1} \frac{F^{0'} F^0 \Lambda^{0'} \tilde{L}_1}{T^2}$ such that $H \Lambda^0 = \frac{1}{NT^2} \Lambda^0 F^{0'} F^0 \Lambda^{0'}$. We have

$$\begin{aligned} & \tilde{l}_{1,i} - H \lambda_{i,0} \\ = & \lambda_{NT}^{-1} \left[\frac{T}{NT^2} \sum_{j=1}^N \tilde{l}_{1,j} E \left(\frac{e'_i e_j}{T} \right) + \frac{T}{NT^2} \sum_{j=1}^N \tilde{l}_{1,j} \left(\frac{e'_i e_j}{T} - \frac{E(e'_i e_j)}{T} \right) \right. \\ & \left. \frac{1}{NT^2} \lambda_{0,i} \sum_{j=1}^N \tilde{l}_{1,j} F^{0'} e_j + \frac{1}{NT^2} e'_i F^0 \sum_{j=1}^N \lambda_{0,j} \tilde{l}_{1,j} \right] \\ = & \lambda_{NT}^{-1} \left[\frac{1}{NT} \sum_{j=1}^N \tilde{l}_{1,j} \gamma_T(i, j) + \frac{1}{NT} \sum_{j=1}^N \tilde{l}_{1,j} \zeta_{ij} + \frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j} \eta_{ij} + \frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j} \xi_{ij} \right], \end{aligned} \quad (5)$$

where $\gamma_T(i, j) = E(e'_i e_j) / T$, $\zeta_{ij} = e'_i e_j / T - \gamma_T(i, j)$, $\eta_{ij} = \lambda_{0,i} F^{0'} e_j / T^2$, and $\xi_{ij} = e'_i F^0 \lambda_{0,j} / T^2$.

A.1 Auxiliary Lemmas

Lemma S.1, S.2 and S.3 are given here to aid the proof of Lemma 4.1.

Lemma S.1. *Under the data generating process (1) and Assumptions 4.1-4.5, as $N, T \rightarrow \infty$, we have*

- (1) $\frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N \gamma_T(i, j)^2 \leq M$;
- (2) $\lambda_{NT} = O_p(1)$;
- (3) $H = \lambda_{NT}^{-1} \frac{F^{0'} F^0 \Lambda^{0'} \tilde{L}_1}{T^2} = O_p(1)$;

Proof . (1) *The proof is as in Chen, Li & Phillips (2019) and is given here for convenience. Let $\rho_{i,j} = \gamma_T(i, j) / [\gamma_T(i, i) \gamma_T(j, j)]^{1/2}$ and by construction $\rho_{i,j}^2 \leq |\rho_{i,j}| \leq 1$. Since $\gamma_T(i, i) \leq M$,*

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N \gamma_T(i, j)^2 &= \frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N \rho_{i,j}^2 \gamma_T(i, i) \gamma_T(j, j) \\ &\leq M \frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N |\rho_{i,j}| |\gamma_T(i, i) \gamma_T(j, j)|^{1/2} \\ &\leq M \frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N |\gamma_T(i, j)| \leq M^2 \end{aligned}$$

by Assumption 4.3(2).

(2) λ_{NT} denotes the leading eigenvalue of $\frac{1}{NT^2} X' X$ and is $O_p(1)$. The result is from Bai (2004, Lemma B.3) and Lemma A.3 in Chen, Li & Phillips (2019). A different proof is given here. From the definition of the eigenvalues and eigenvectors and the normalization $\tilde{L}'_1 \tilde{L}_1 / N = 1$, we have

$$\begin{aligned} \lambda_{NT} &= \frac{1}{N^2 T^2} \tilde{L}'_1 X' X \tilde{L}_1 \\ &= \frac{1}{N^2 T^2} \tilde{L}'_1 (\Lambda^0 F^{0'} F^0 \Lambda^{0'} + \Lambda^0 F^{0'} E + E' F^0 \Lambda^{0'} + E' E) \tilde{L}_1. \end{aligned}$$

The first term

$$\frac{1}{N^2 T^2} \tilde{L}'_1 \Lambda^0 F^{0'} F^0 \Lambda^0 \tilde{L}_1 = \frac{1}{N^2} \left(\frac{F^{0'} F^0}{T^2} \right) \left(\sum_{i=1}^N l_{1,i} \lambda_{0,i} \right)^2 = O_p(1),$$

is from Lemma A.2 and by Cauchy-Schwartz inequality, we obtain

$$\sum_{i=1}^N \tilde{l}_{1,i} \lambda_{0,i} \leq N \left(\frac{1}{N} \sum_{i=1}^N \tilde{l}_{1,i}^2 \right)^{1/2} \left(\frac{1}{N} \sum_{i=1}^N \lambda_{0,i}^2 \right)^{1/2} = O_p(N). \quad (6)$$

using Assumption 4.2(1). The second term

$$\begin{aligned} \frac{1}{N^2 T^2} \tilde{L}'_1 \Lambda^0 F^{0'} E \tilde{L}_1 &= \frac{1}{N^2 T^2} \left(\sum_{i=1}^N \tilde{l}_{1,i} \lambda_{0,i} \right) \left(\sum_{t=1}^T \sum_{i=1}^N \tilde{l}_{1,i} f_{0,t} e_{it} \right) \\ &= \frac{1}{N^2 T} \left(\sum_{i=1}^N \tilde{l}_{1,i} \lambda_{0,i} \right) \sum_{i=1}^N \tilde{l}_{1,i} \left(\frac{1}{T} \sum_{t=1}^T f_{0,t} e_{it} \right) \\ &= O_p \left(\frac{1}{T} \right) \end{aligned}$$

from Assumption 4.4, equation (6), and

$$\left(\sum_{i=1}^N \tilde{l}_{1,i} \right)^2 \leq \left(\sum_{i=1}^N \tilde{l}_{1,i}^2 \right) \left(\sum_{i=1}^N 1 \right) = N \sum_{i=1}^N \tilde{l}_{1,i}^2 = N^2 \left(\frac{1}{N} \sum_{i=1}^N \tilde{l}_{1,i}^2 \right) = O_p(N^2). \quad (7)$$

by the Cauchy-Schwartz inequality. The third term is $O_p\left(\frac{1}{T}\right)$ since

$$\frac{1}{N^2 T^2} \tilde{L}'_1 E' F^0 \Lambda^0 \tilde{L}_1 = \frac{1}{N^2 T^2} \left(\sum_{t=1}^T \sum_{i=1}^N \tilde{l}_{1,i} f_{0,t} e_{it} \right) \left(\sum_{i=1}^N \tilde{l}_{1,i} \lambda_{0,i} \right) = O_p \left(\frac{1}{T} \right).$$

The fourth term is $O_p\left(\frac{1}{T}\right)$ since

$$\frac{1}{N^2 T^2} \tilde{L}'_1 E' E \tilde{L}_1 = \frac{1}{N^2 T^2} \sum_{t=1}^T \left(\sum_{i=1}^N \tilde{l}_{1,i} e_{it} \right)^2 \leq \frac{1}{T^2} \sum_{t=1}^T \left(\frac{1}{N} \sum_{i=1}^N \tilde{l}_{1,i}^2 \right) \left(\frac{1}{N} \sum_{i=1}^N e_{it}^2 \right) = O_p \left(\frac{1}{T} \right).$$

Thus, the first term dominates the others and we establish

$$\lambda_{NT} = \frac{1}{N^2 T^2} \tilde{L}'_1 \Lambda^0 F^{0'} F^0 \Lambda^0 \tilde{L}_1 [1 + o_p(1)] = O_p(1).$$

(3) The normalization variable

$$H = \lambda_{NT}^{-1} \frac{F^{0'} F^0}{T^2} \frac{\Lambda^0 \tilde{L}_1}{N} = \frac{1}{N} \lambda_{NT}^{-1} \frac{F^{0'} F^0}{T^2} \left(\sum_{i=1}^N \tilde{l}_{1,i} \lambda_{0,i} \right) = O_p(1),$$

from Lemma S1(2), Lemma A.2(1) and (6).

Lemma S.2. Under the data generating process (1) and Assumptions 4.1-4.4, as $N, T \rightarrow \infty$, we have

$$\frac{1}{N} \sum_{i=1}^N (\tilde{l}_{1,i} - H \lambda_{0,i})^2 = O_p(T^{-2}).$$

Proof . (1) Using (5) and the inequality $(x+y+z+u)^2 \leq 4(x^2+y^2+z^2+u^2)$, we have

$$(\tilde{l}_{1,i} - H\lambda_{i,0})^2 \leq 4(a_i^2 + b_i^2 + c_i^2 + d_i^2),$$

where

$$\begin{aligned} a_i^2 &= \lambda_{NT}^{-2} \frac{1}{N^2 T^2} \left(\sum_{j=1}^N \tilde{l}_{1,j} \gamma_T(i,j) \right)^2, \quad b_i^2 = \lambda_{NT}^{-2} \frac{1}{N^2 T^2} \left(\sum_{j=1}^N \tilde{l}_{1,j} \zeta_{ij} \right)^2, \\ c_i^2 &= \lambda_{NT}^{-2} \frac{1}{N^2} \left(\sum_{j=1}^N \tilde{l}_{1,j} \eta_{ij} \right)^2, \quad \text{and } d_i^2 = \lambda_{NT}^{-2} \frac{1}{N^2} \left(\sum_{j=1}^N \tilde{l}_{1,j} \xi_{ij} \right)^2. \end{aligned}$$

It follows that

$$\frac{1}{N} \sum_{i=1}^N (\tilde{l}_{1,i} - H\lambda_{i,0})^2 \leq 4 \left(\frac{1}{N} \sum_{i=1}^N a_i^2 + \frac{1}{N} \sum_{i=1}^N b_i^2 + \frac{1}{N} \sum_{i=1}^N c_i^2 + \frac{1}{N} \sum_{i=1}^N d_i^2 \right).$$

The first term, by the Cauchy-Schwarz inequality, is

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N a_i^2 &= \lambda_{NT}^{-2} \frac{1}{N^3 T^2} \sum_{i=1}^N \left(\sum_{j=1}^N \tilde{l}_{1,j} \gamma_T(i,j) \right)^2 \\ &\leq \frac{1}{NT^2} \lambda_{NT}^{-2} \sum_{i=1}^N \left(\frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j}^2 \right) \left(\frac{1}{N} \sum_{j=1}^N \gamma_T(i,j)^2 \right) \\ &= \frac{1}{NT^2} \lambda_{NT}^{-2} \left(\frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j}^2 \right) \left(\frac{1}{N} \sum_{i=1}^N \sum_{j=1}^N \gamma_T(i,j)^2 \right) \\ &= O_p(N^{-1} T^{-2}), \end{aligned}$$

from Lemma S.1(1) and Lemma S.1(2) and the normalization $\frac{1}{N} \tilde{L}'_1 \tilde{L}_1 = 1$. The second term is

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N b_i^2 &= \lambda_{NT}^{-2} \frac{1}{N^3 T^2} \sum_{i=1}^N \left(\sum_{j=1}^N \tilde{l}_{1,j} \zeta_{ij} \right)^2 \\ &= \lambda_{NT}^{-2} \frac{1}{N^3 T^2} \sum_{i=1}^N \sum_{j,u=1}^N \tilde{l}_{1,j} \tilde{l}_{1,u} \zeta_{ij} \zeta_{iu} \\ &= \lambda_{NT}^{-2} \frac{1}{N^3 T^2} \sum_{j,u=1}^N (\tilde{l}_{1,j} \tilde{l}_{1,u}) \left(\sum_{i=1}^N \zeta_{ij} \zeta_{iu} \right) \\ &\leq \lambda_{NT}^{-2} \frac{1}{N^3 T^2} \left[\sum_{j,u=1}^N (\tilde{l}_{1,j} \tilde{l}_{1,u})^2 \right]^{1/2} \left[\sum_{j,u=1}^N \left(\sum_{i=1}^N \zeta_{ij} \zeta_{iu} \right)^2 \right]^{1/2} \\ &\leq \lambda_{NT}^{-2} \frac{1}{NT^2} \left(\frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j}^2 \right) \left[\frac{1}{N^2} \sum_{j,u=1}^N \left(\sum_{i=1}^N \zeta_{ij} \zeta_{iu} \right)^2 \right]^{1/2} \\ &= \lambda_{NT}^{-2} \frac{1}{NT^2} \left[\frac{1}{N^2} \sum_{j,u=1}^N \left(\sum_{i=1}^N \zeta_{ij} \zeta_{iu} \right)^2 \right]^{1/2}. \end{aligned}$$

We know that

$$E \left(\sum_{i=1}^N \zeta_{ij} \zeta_{iu} \right)^2 = E \left(\sum_{i,s=1}^N \zeta_{ij} \zeta_{iu} \zeta_{sj} \zeta_{su} \right) \leq N^2 \max E |\zeta_{ij}|^4,$$

and

$$E |\zeta_{ij}|^4 = E \left| \frac{\mathbf{e}'_i \mathbf{e}_j}{T} - \gamma_T(i, j) \right|^4 = \left(\frac{1}{T^2} \right) E \left| T^{1/2} \left(\frac{\mathbf{e}'_i \mathbf{e}_j}{T} - \gamma_T(i, j) \right) \right|^4 \leq \frac{1}{T^2} M,$$

from Assumption 4.3 (5). Thus,

$$\frac{1}{N} \sum_{i=1}^N b_i^2 = O_p(T^{-3}).$$

The third term is

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N c_i^2 &= \lambda_{NT}^{-2} \frac{1}{N^3} \sum_{i=1}^N \left(\sum_{j=1}^N \tilde{l}_{1,j} \eta_{ij} \right)^2 \\ &\leq \lambda_{NT}^{-2} \frac{1}{N^2} \sum_{i=1}^N \left(\frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j}^2 \right) \left(\sum_{j=1}^N \eta_{ij}^2 \right) \\ &= \lambda_{NT}^{-2} \frac{1}{N} \left[\frac{1}{T^2} \sum_{j=1}^N \left(\frac{1}{T} F^{0r} \mathbf{e}_j \right)^2 \right] \left(\frac{1}{N} \sum_{i=1}^N \lambda_{0,i}^2 \right) \\ &= O_p(T^{-2}), \end{aligned}$$

from Assumption 4.2(1) and Lemma S.1 (1) and by virtue of Assumption 4 since

$$\left[\frac{1}{N} \sum_{j=1}^N \left(\frac{1}{T} F^{0r} \mathbf{e}_j \right)^2 \right] \leq \left(\sup_{j \geq 1} \left| \frac{1}{T} F^{0r} \mathbf{e}_j \right| \right)^2 = O_p(1).$$

Similarly, the last term is

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N d_i^2 &= \lambda_{NT}^{-2} \frac{1}{N^3} \sum_{i=1}^N \left(\sum_{j=1}^N \tilde{l}_{1,j} \xi_{ij} \right)^2 \\ &= \lambda_{NT}^{-2} \frac{1}{N^3} \frac{1}{T^4} \left(\sum_{j=1}^N \tilde{l}_{1,j} \lambda_{0,j} \right)^2 \left(\sum_{i=1}^N (\mathbf{e}'_i F^0)^2 \right) \\ &= O_p(T^{-2}), \end{aligned}$$

from (6) and Assumption 4.4. Therefore, the third and fourth terms dominate and

$$\frac{1}{N} \sum_{i=1}^N (\tilde{l}_{1,i} - H \lambda_{0,i})^2 = O_p(T^{-2}).$$

Lemma S.3. Under the data generation (1) and Assumption 4.1-4.5, as $T, N \rightarrow \infty$,

$$\Lambda^{0r} (\Lambda^0 H - \tilde{L}_1) / N = O_p \left(\frac{1}{T \sqrt{N}} \right).$$

Proof. (i) Using (5), we have

$$\begin{aligned} &\Lambda^{0r} (\Lambda^0 H - \tilde{L}_1) / N \\ &= -\frac{1}{N} \sum_{i=1}^N (\tilde{l}_{1,i} - H \lambda_{0,i}) \lambda_{0,i} \\ &= -\frac{1}{N} \lambda_{NT}^{-1} \sum_{i=1}^N \left[\frac{1}{NT^2} \sum_{j=1}^N \tilde{l}_{1,j} \mathbf{e}'_i \mathbf{e}_j + \frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j} \eta_{ij} + \frac{1}{N} \sum_{j=1}^N \tilde{l}_{1,j} \xi_{ij} \right] \lambda_{0,i} \\ &= -H \lambda_{NT}^{-1} \left[\frac{1}{N^2 T^2} \sum_{i=1}^N \sum_{j=1}^N \lambda_{0,j} \mathbf{e}'_i \mathbf{e}_j \lambda_{0,i} + \frac{1}{N^2} \sum_{i=1}^N \sum_{j=1}^N \lambda_{0,j} \eta_{ij} \lambda_{0,i} + \frac{1}{N^2} \sum_{i=1}^N \sum_{j=1}^N \lambda_{0,j} \xi_{ij} \lambda_{0,i} \right] [1 + o_p(1)] \\ &= I + II + III \end{aligned}$$

where I, II, III correspond to the three component sums. The second to last equality is due to the fact that

$$\tilde{l}_{1,j} = (\tilde{l}_{1,j} - H\lambda_{0,j}) + H\lambda_{0,j} = H\lambda_{0,j}[1 + o_p(1)],$$

which follows directly from the proof of Lemma S.2. The first term I is

$$-H\lambda_{NT}^{-1} \frac{1}{N^2 T^2} \sum_{i=1}^N \lambda_{0,i} e_i' \sum_{j=1}^N \lambda_{0,j} e_j = -\frac{1}{NT} H\lambda_{NT}^{-1} \left[\frac{1}{T} \sum_{t=1}^T \left(\frac{1}{\sqrt{N}} \sum_{j=1}^N \lambda_{0,j} e_{jt} \right)^2 \right] = O_p(N^{-1} T^{-1}),$$

by Lemma 1(ii) in Bai and Ng (2002). The second term II is

$$-\lambda_{NT}^{-1} \frac{1}{N^2} \sum_{i=1}^N \sum_{j=1}^N H\lambda_{0,j} \eta_{ij} \lambda_{0,i} = -\frac{1}{NT^2} H\lambda_{NT}^{-1} \left(\frac{1}{N} \sum_{i=1}^N \lambda_{0,i}^2 \right) \sum_{t=1}^T \sum_{j=1}^N \lambda_{0,j} f_{0,t} e_{jt} = O_p\left(\frac{1}{\sqrt{NT}}\right),$$

since

$$\sum_{t=1}^T \sum_{j=1}^N \lambda_{0,j} f_{0,t} e_{jt} = T \sum_{i=1}^N \lambda_{0,i} \left(\frac{1}{T} \sum_{t=1}^T f_{0,t} e_{it} \right) = O_p(T\sqrt{N}).$$

The third term III is

$$-\lambda_{NT}^{-1} \frac{1}{N^2} \sum_{i=1}^N \sum_{j=1}^N H\lambda_{0,j} \xi_{ij} \lambda_{0,i} = -H\lambda_{NT}^{-1} \frac{1}{NT^2} \left(\frac{1}{N} \sum_{j=1}^N \lambda_{0,j}^2 \right) \sum_{i=1}^N e_i' F^0 \lambda_{0,i} = O_p\left(\frac{1}{\sqrt{NT}}\right).$$

Thus, we have

$$\Lambda^{0r} (\Lambda^0 H - \tilde{L}_1) / N = I + II + III = O_p\left(\frac{1}{\sqrt{NT}}\right).$$

Lemma 4.1. Under Assumptions 4.1-4.5, as $N, T \rightarrow \infty$, we have

$$\delta_{NT}^2 \left(\frac{1}{T} \sum_{t=1}^T |\tilde{y}_t - H^0 f_{0,t}|^2 \right) = O_p(1), \quad (8)$$

where $\delta_{NT} = \min(\sqrt{N}, T)$ and $H^0 = H^{-1}$.

Proof. By construction,

$$\tilde{y}_t = \tilde{L}'_1 X_t / N = \tilde{L}'_1 (\Lambda^0 f_{0,t} + e_t) / N.$$

Rewriting $\Lambda^0 = \Lambda^0 - \tilde{L}_1 H^{-1} + \tilde{L}_1 H^{-1}$ and using the restriction $\tilde{L}'_1 \tilde{L}_1 / N = 1$, we have

$$\tilde{y}_t = H^{-1} f_{0,t} + \tilde{L}'_1 (\Lambda^0 - \tilde{L}_1 H^{-1}) f_{0,t} / N + (\tilde{L}'_1 - H\Lambda^{0r}) e_t / N + H\Lambda^{0r} e_t / N$$

As before, we use the inequality

$$(\tilde{y}_t - H^{-1} f_{0,t})^2 \leq 4(a_t^2 + b_t^2 + c_t^2),$$

where

$$a_t^2 = \frac{1}{N^2} [\tilde{L}'_1 (\Lambda^0 - \tilde{L}_1 H^{-1}) f_{0,t}]^2, b_t^2 = \frac{1}{N^2} [(\tilde{L}'_1 - H\Lambda^{0r}) e_t]^2, \text{ and } c_t^2 = \frac{1}{N^2} (H\Lambda^{0r} e_t)^2.$$

It follows that

$$\frac{1}{T} \sum_{t=1}^T (\tilde{y}_t - H^{-1} f_{0,t})^2 \leq 4 \left(\frac{1}{T} \sum_{t=1}^T a_t^2 + \frac{1}{T} \sum_{t=1}^T b_t^2 + \frac{1}{T} \sum_{t=1}^T c_t^2 \right).$$

Rewrite $\tilde{L}'_1 (\Lambda^0 H - \tilde{L})$ as

$$\begin{aligned}\tilde{L}'_1 (\Lambda^0 H - \tilde{L}_1) / N &= (\tilde{L}_1 - \Lambda^0 H)' (\Lambda^0 H - \tilde{L}_1) / N + H \Lambda^{0'} (\Lambda^0 H - \tilde{L}_1) / N \\ &= O_p \left(\frac{1}{T^2} \right) + O_p \left(\frac{1}{T\sqrt{N}} \right) = O_p \left(\frac{1}{T\sqrt{N}} \right)\end{aligned}$$

from Lemma S.2 and Lemma S.3, respectively. The first term is

$$\begin{aligned}\frac{1}{T} \sum_{t=1}^T a_t^2 &= \frac{1}{T} \sum_{t=1}^T [\tilde{L}'_1 (\Lambda^0 H - \tilde{L}_1) / N]^2 H^{-2} f_{0,t}^2 \\ &= \frac{1}{T} H^{-2} [\tilde{L}'_1 (\Lambda^0 H - \tilde{L}_1) / N]^2 \sum_{t=1}^T f_{0,t}^2 \\ &= O_p(T^{-1}) O_p(1) O_p \left(\frac{1}{T^2 N} \right) O_p(T^2) = O_p \left(\frac{1}{TN} \right).\end{aligned}$$

The second term is

$$\begin{aligned}\frac{1}{T} \sum_{t=1}^T b_t^2 &= \frac{1}{N^2 T} \sum_{t=1}^T [(\tilde{L}_1 - H \Lambda^{0'}) e_t]^2 = \frac{1}{N^2 T} \sum_{t=1}^T \left(\sum_{i=1}^N (\tilde{l}_{1,i} - H \lambda_{0,i}) e_{it} \right)^2 \\ &\leq \frac{1}{T} \sum_{t=1}^T \left(\frac{1}{N} \sum_{i=1}^N e_{it}^2 \right) \left(\frac{1}{N} \sum_{i=1}^N (\tilde{l}_{1,i} - H \lambda_{0,i})^2 \right) \\ &= \left(\frac{1}{N} \sum_{i=1}^N (\tilde{l}_{1,i} - H \lambda_{0,i})^2 \right) \frac{1}{NT} \sum_{t=1}^T \sum_{i=1}^N e_{it}^2 = O_p(T^{-2}),\end{aligned}$$

from Lemma A.1 and $\frac{1}{NT} \sum_{t=1}^T \sum_{i=1}^N e_{it}^2 = O_p(1)$ from the Assumption 4.3(4). The third term is

$$\frac{1}{T} \sum_{t=1}^T c_t^2 = \frac{1}{TN^2} H^2 \sum_{t=1}^T (\Lambda^{0'} e_t)^2 = \frac{1}{N} H^2 \frac{1}{T} \sum_{t=1}^T \left(\frac{1}{N^{1/2}} \sum_{i=1}^N \lambda_{0,i} e_{it} \right)^2 = O_p(N^{-1}),$$

by Lemma 1(ii) in Bai and Ng (2002) which requires Assumption 4.2(1) and Assumption 4.3(3). Therefore, we have

$$\frac{1}{T} \sum_{t=1}^T (\tilde{y}_t - H^{-1} f_{0,t})^2 = O_p \left(\frac{1}{T^2} \right) + O_p \left(\frac{1}{N} \right) = O_p(\delta_{NT}^{-2}),$$

where $\delta_{NT} = \min(\sqrt{N}, T)$.

B Finite Sample Performance of the Real-time PSY-factor Procedure

This section presents Monte Carlo simulation results for the real-time PSY-factor procedure. Instead of estimating the first factor from the entire sample, we compute the factor from a sample starting with the first available observation and ending with the current observation at τ using only historical information up to this point in time. Figures 1 and 2 report empirical sizes, SDRs and estimation accuracy of the bubble origination date by the real time recursive procedure. We set $\sigma_{00} = 0.08$ and $\sigma_{11} = 0.1$. Results for the case of $\sigma_{00} = 0.02$ and $\sigma_{11} = 0.06$ were found to be similar and are not reported. The performance of the real-time procedure is almost identical to results in Section 5 of the main paper for the PSY-factor procedure.

Figure 1: Empirical sizes of the real-time PSY-factor procedure. The nominal size is 5%.

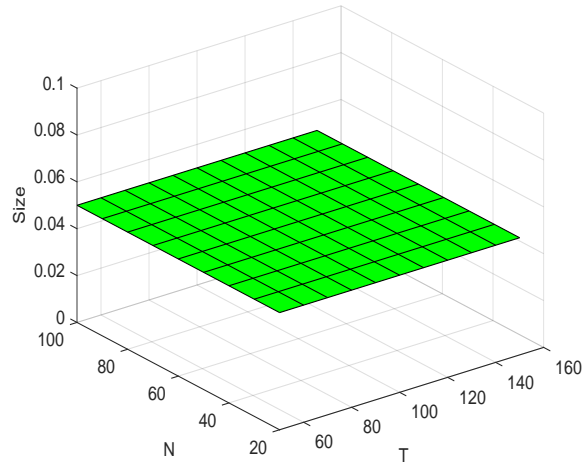
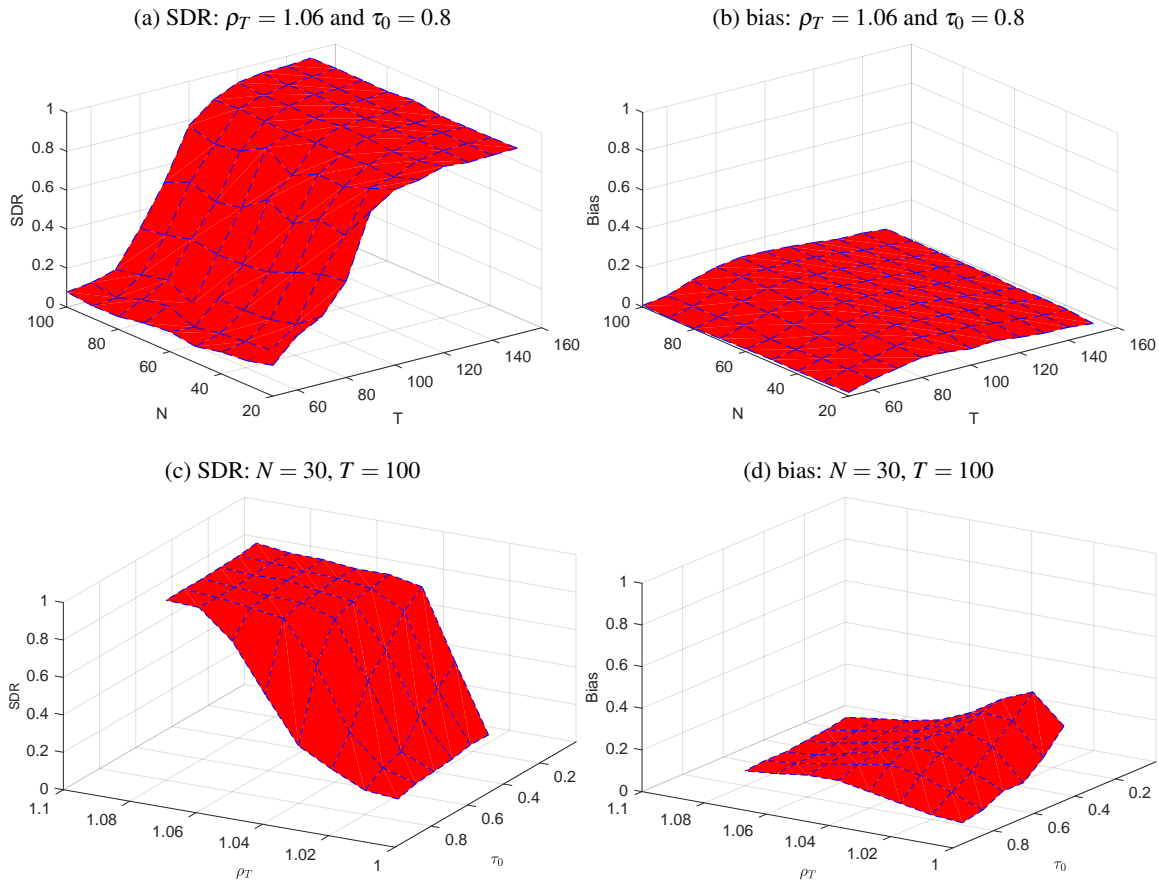


Figure 2: The successful detection rates (SDRs) and bias of the estimated bubble origination of the real time PSY-factor procedure. Parameter settings: $\tau_0 = 0.8$, $r_0 = 0.7$ and $\rho_T = 1.06$, $\sigma_{00} = 0.08$ and $\sigma_{11} = 0.1$.



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